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## Great Expectations

Earlier in the course we came to the important result that $E\left[\sum_{i} X_{i}\right]=\sum_{i} E\left[X_{i}\right]$. First, as a warm up lets go back to our old friends and show how we could have derived expressions for their expectation.

## Expectation of Binomial

First let's start with some practice with the sum of expectations of indicator variables. Let $Y \sim \operatorname{Bin}(n, p)$, in other words if $Y$ is a Binomial random variable. We can express $Y$ as the sum of $n$ Bernoulli random indicator variables $X_{i} \sim \operatorname{Ber}(p)$. Since $X_{i}$ is a Bernoulli, $E\left[X_{i}\right]=p$

$$
Y=X_{1}+X_{2}+\cdots+X_{n}=\sum_{i=1}^{n} X_{i}
$$

Let's formally calculate the expectation of $Y$ :

$$
\begin{aligned}
E[Y] & =E\left[\sum_{i}^{n} X_{i}\right] \\
& =\sum_{i}^{n} E\left[X_{i}\right] \\
& =E\left[X_{0}\right]+E\left[X_{1}\right]+\ldots E\left[X_{n}\right] \\
& =n p
\end{aligned}
$$

## Expectation of Negative Binomial

Recall that a Negative Binomial is a random variable that semantically represents the number of trials until $r$ successes. Let $Y \sim N e g \operatorname{Bin}(r, p)$.

Let $X_{i}=\#$ trials to get success after $(i-1)$ st success. We can then think of each $X_{i}$ as a Geometric RV: $X_{i} \sim \operatorname{Geo}(p)$. Thus, $E\left[X_{i}\right]=\frac{1}{p}$. We can express $Y$ as:

$$
Y=X_{1}+X_{2}+\cdots+X_{r}=\sum_{i=1}^{r} X_{i}
$$

Let's formally calculate the expectation of $Y$ :

$$
\begin{aligned}
E[Y] & =E\left[\sum_{i=1}^{r} X_{i}\right] \\
& =\sum_{i=1}^{r} E\left[X_{i}\right] \\
& =E\left[X_{1}\right]+E\left[X_{2}\right]+\ldots E\left[X_{r}\right] \\
& =\frac{r}{p}
\end{aligned}
$$

## Conditional Expectation

We have gotten to know a kind and gentle soul, conditional probability. And we now know another funky fool, expectation. Let's get those two crazy kids to play together.

Let $X$ and $Y$ be jointly random variables. Recall that the conditional probability mass function (if they are discrete), and the probability density function (if they are continuous) are respectively:

$$
\begin{aligned}
p_{X \mid Y}(x \mid y) & =\frac{p_{X, Y}(x, y)}{p_{Y}(y)} \\
f_{X \mid Y}(x \mid y) & =\frac{f_{X, Y}(x, y)}{f_{Y}(y)}
\end{aligned}
$$

We define the conditional expectation of $X$ given $Y=y$ to be:

$$
\begin{aligned}
& E[X \mid Y=y]=\sum_{x} x p_{X \mid Y}(x \mid y) \\
& E[X \mid Y=y]=\int_{-\infty}^{\infty} x f_{X \mid Y}(x \mid y) d x
\end{aligned}
$$

Where the first equation applies if $X$ and $Y$ are discrete and the second applies if they are continuous.

## Properties of Conditional Expectation

Here are some helpful, intuitive properties of conditional expectation:

$$
\begin{aligned}
E[g(X) \mid Y=y] & =\sum_{x} g(x) p_{X \mid Y}(x \mid y) & \text { if } \mathrm{X} \text { and Y are discrete } \\
E[g(X) \mid Y=y] & =\int_{-\infty}^{\infty} g(x) f_{X \mid Y}(x \mid y) d x & \text { if } \mathrm{X} \text { and } \mathrm{Y} \text { are continuous } \\
E\left[\sum_{i=1}^{n} X_{i} \mid Y=y\right] & =\sum_{i=1}^{n} E\left[X_{i} \mid Y=y\right] &
\end{aligned}
$$

## Law of Total Expectation

The law of total expectation states that: $\mathrm{E}[\mathrm{E}[\mathrm{X}-\mathrm{Y}]]=\mathrm{E}[\mathrm{X}]$
What?! How is that a thing? Check out this proof:

$$
\begin{aligned}
E[E[X \mid Y]] & =\sum_{y} E[X \mid Y=y] P(Y=y) \\
& =\sum_{y} \sum_{x} x P(X=x \mid Y=y) P(Y=y) \\
& =\sum_{y} \sum_{x} x P(X=x, Y=y) \\
& =\sum_{x} \sum_{y} x P(X=x, Y=y) \\
& =\sum_{x} x \sum_{y} P(X=x, Y=y) \\
& =\sum_{x} x P(X=x) \\
& =E[X]
\end{aligned}
$$

## Example 1

You roll two 6-sided dice $D_{1}$ and $D_{2}$. Let $X=D_{1}+D_{2}$ and let $Y=$ the value of $D_{2}$. What is $E[X \mid Y=6]$

$$
\begin{aligned}
E[X \mid Y=6] & =\sum_{x} x P(X=x \mid Y=6) \\
& =\left(\frac{1}{6}\right)(7+8+9+10+11+12)=\frac{57}{6}=9.5
\end{aligned}
$$

Which makes intuitive sense since $6+E\left[\right.$ value of $\left.D_{1}\right]=6+3.5$

## Example 2

Consider the following code with random numbers:

```
int Recurse() {
    int x = randomInt(1, 3); // Equally likely values
    if (x == 1) return 3;
    else if (x == 2) return (5 + Recurse());
    else return (7 + Recurse());
}
```

Let $Y=$ value returned by "Recurse". What is $E[Y]$. In other words, what is the expected return value. Note that this is the exact same approach as calculating the expected run time.

$$
E[Y]=E[Y \mid X=1] P(X=1)+E[Y \mid X=2] P(X=2)+E[Y \mid X=3] P(X=3)
$$

First lets calculate each of the conditional expectations:

$$
\begin{aligned}
& E[Y \mid X=1]=3 \\
& E[Y \mid X=2]=E[5+Y]=5+E[Y] \\
& E[Y \mid X=3]=E[7+Y]=7+E[Y]
\end{aligned}
$$

Now we can plug those values into the equation. Note that the probability of $X$ taking on 1,2 , or 3 is $1 / 3$ :

$$
\begin{aligned}
E[Y] & =E[Y \mid X=1] P(X=1)+E[Y \mid X=2] P(X=2)+E[Y \mid X=3] P(X=3) \\
& =3(1 / 3)+(5+E[Y])(1 / 3)+(7+E[Y])(1 / 3) \\
& =15
\end{aligned}
$$

## Hiring Software Engineers

You are interviewing $n$ software engineer candidates and will hire only 1 candidate. All orderings of candidates are equally likely. Right after each interview you must decide to hire or not hire. You can not go back on a decision. At any point in time you can know the relative ranking of the candidates you have already interviewed.

The strategy that we propose is that we interview the first $k$ candidates and reject them all. Then you hire the next candidate that is better than all of the first $k$ candidates. What is the probability that the best of all the $n$ candidates is hired for a particular choice of $k$ ? Let's denote that result $P_{k}($ Best $)$. Let $X$ be the position in the ordering of the best candidate:

$$
P_{k}(\text { Best })=\sum_{i=1}^{n} P_{k}(\text { Best } \mid X=i) P(X=i)
$$

$$
=\frac{1}{n} \sum_{i=1}^{n} P_{k}(\text { Best } \mid X=i) \quad \text { since each position is equally likely }
$$

What is $P_{k}($ Best $\mid X=i)$ ? if $i \leq k$ then the probability is 0 because the best candidate will be rejected without consideration. Sad times. Otherwise we will chose the best candidate, who is in position $i$, only if the best of the first $i-1$ candidates is among the first $k$ interviewed. If the best among the first $i-1$ is not among the first $k$, that candidate will be chosen over the true best. Since all orderings are equally likely the probability that the best among the $i-1$ candidates is in the first $k$ is:

$$
\frac{k}{i-1}
$$

if $i>k$

Now we can plug this back into our original equation:

$$
\begin{array}{rlr}
P_{k}(\text { Best }) & =\frac{1}{n} \sum_{i=1}^{n} P_{k}(\text { Best } \mid X=i) & \\
& =\frac{1}{n} \sum_{i=k+1}^{n} \frac{k}{i-1} & \\
& \approx \frac{1}{n} \int_{i=k+1}^{n} \frac{k}{i-1} d i & \text { since we know } P_{k}(\text { Best } \mid X=i) \\
& =\left.\frac{k}{n} \ln (i=1)\right|_{k+1} ^{n}=\frac{k}{n} \ln \frac{n-1}{k} \approx \frac{k}{n} \ln \frac{n}{k} &
\end{array}
$$

If we think of $P_{k}($ Best $)=\frac{k}{n} \ln \frac{n}{k}$ as a function of $k$ we can take find the value of $k$ that optimizes it by taking its derivative and setting it equal to 0 . The optimal value of $k$ is $n / e$. Where $e$ is Euler's number.

## Lovely lemma

This is not something you will need to use in CS109, but I wanted to show you a neat proof related to expectations. Did you know that if $Y$ is a non-negative random variable the following hold (for discrete and continuous random variables respectively):

$$
\begin{aligned}
& E[Y]=\sum_{i=1}^{n} P(Y \geq i) \\
& E[Y]=\int_{0}^{\infty} P(Y \geq i) d i=\int_{0}^{\infty} F(i) d i
\end{aligned}
$$

Let's explore the lemma. Say $Y$ is a discrete random variable that can take on values in the range $[0, n]$ :

$$
\begin{aligned}
E[Y] & =\sum_{i=1}^{n} P(Y \geq i) & & \text { Why is this true? } \\
& =\sum_{i=1}^{n} \sum_{j=i}^{n} P(Y=j) & & \text { Since the cases of } P(Y \geq i) \text { are mutually exclusive } \\
& =\sum_{k=1}^{n} k \cdot P(Y=k) & & \text { Thinking hard about the double sum. See below }
\end{aligned}
$$

Happily this is the definition of expectation. However how did we simplify that double sum? Think about the double sum: how many times does $j$ take on the value 1 ? Only once (when $i=1$ ). How many times does $j$ take on the value 2? Twice (once when $i=1$ and once when $i=2$ ). In general $j$ takes on the value $k, k$ times.

